

Hands-On Learning Theory

Fall 2016, Lecture 8

Jean Honorio jhonorio@purdue.edu

1 Restricted Strong Convexity

Let \mathbf{w} be a vector and ℓ be a loss function. In general, ℓ_1 -norm regularized loss minimization can be written as follows for some $\lambda > 0$:

$$\hat{\mathbf{w}} = \arg \min_{\mathbf{w} \in \mathbb{R}^p} \ell(\mathbf{w}) + \lambda \|\mathbf{w}\|_1 \quad (1)$$

We will also assume that there is an unknown but fixed $\mathbf{w}^* \in \mathbb{R}^p$. Our goal will be to recover a vector $\hat{\mathbf{w}}$ which is *close to* \mathbf{w}^* .

Next, we define restricted strong convexity [1].

Definition 8.1. Let $\alpha > 0$, $\tau \geq 0$ and $g : \mathbb{N} \times \mathbb{N} \rightarrow \mathbb{R}^+$. A loss function $\ell : \mathbb{R}^p \rightarrow \mathbb{R}$ is *restricted strongly convex around \mathbf{w}^** with parameters (α, τ, g) if and only if:

$$(\forall \mathbf{w} \in \mathbb{R}^p) \ell(\mathbf{w}) - \ell(\mathbf{w}^*) - \langle \nabla \ell(\mathbf{w}^*), \mathbf{w} - \mathbf{w}^* \rangle \geq \alpha \|\mathbf{w} - \mathbf{w}^*\|_2^2 - \tau g(n, p) \|\mathbf{w} - \mathbf{w}^*\|_1^2$$

For many specific learning problems $g(n, p) = \sqrt{\frac{\log p}{n}}$ or $g(n, p) = \frac{\log p}{n}$. In what follows, we analyze the sufficient number of samples for the problem in eq.(1).

Theorem 8.1. Assume that the convex loss function $\ell : \mathbb{R}^p \rightarrow \mathbb{R}$ is *restricted strongly convex around \mathbf{w}^** with parameters (α, τ, g) as in Definition 8.1. Let k be the number of nonzero elements in \mathbf{w}^* . For a regularization weight $\lambda \geq 2\|\nabla \ell(\mathbf{w}^*)\|_\infty$ and a sufficient number of samples $17\frac{\tau}{\alpha}kg(n, p) \leq 1$, we have that:

$$\|\hat{\mathbf{w}} - \mathbf{w}^*\|_2 \leq 102 \frac{\sqrt{k}}{\alpha} \lambda$$
$$\|\hat{\mathbf{w}} - \mathbf{w}^*\|_1 \leq 408 \frac{k}{\alpha} \lambda$$

For many specific learning problems $\lambda \in \mathcal{O}(\sqrt{\frac{\log p}{n}})$ and thus, the above theorem establishes consistency as the number of samples n grows.

First, we derive an intermediate lemma needed for the final proof.

Lemma 8.1. *Assume that the loss function $\ell : \mathbb{R}^p \rightarrow \mathbb{R}$ is convex. Let k be the number of nonzero elements in \mathbf{w}^* . For a regularization weight $\lambda \geq 2\|\nabla\ell(\mathbf{w}^*)\|_\infty$, we have:*

$$\|\widehat{\mathbf{w}} - \mathbf{w}^*\|_1 \leq 4\sqrt{k}\|\widehat{\mathbf{w}} - \mathbf{w}^*\|_2$$

Proof. Let $\Delta \equiv \widehat{\mathbf{w}} - \mathbf{w}^*$. Let \mathcal{K} be the set of nonzero elements of \mathbf{w}^* and let \mathcal{K}^c be the complement of \mathcal{K} . Note that $k \equiv |\mathcal{K}|$ is the number of nonzero elements in \mathbf{w}^* . For an arbitrary vector \mathbf{w} , let $\mathbf{w}_{\mathcal{K}}$ denote the original vector \mathbf{w} with zeros on the entries in \mathcal{K}^c . Similarly, let $\mathbf{w}_{\mathcal{K}^c}$ denote the original vector \mathbf{w} with zeros on the entries in \mathcal{K} .

Since by definition $\mathbf{w}^* = \mathbf{w}_{\mathcal{K}}^*$ and by the reverse triangle inequality, we have:

$$\begin{aligned} \|\widehat{\mathbf{w}}\|_1 &= \|\mathbf{w}^* + \Delta\|_1 \\ &= \|\mathbf{w}_{\mathcal{K}}^* + \Delta_{\mathcal{K}} + \Delta_{\mathcal{K}^c}\|_1 \\ &= \|\mathbf{w}_{\mathcal{K}}^* + \Delta_{\mathcal{K}}\|_1 + \|\Delta_{\mathcal{K}^c}\|_1 \\ &\geq \|\mathbf{w}_{\mathcal{K}}^*\|_1 - \|\Delta_{\mathcal{K}}\|_1 + \|\Delta_{\mathcal{K}^c}\|_1 \\ &= \|\mathbf{w}^*\|_1 - \|\Delta_{\mathcal{K}}\|_1 + \|\Delta_{\mathcal{K}^c}\|_1 \end{aligned} \quad (2)$$

By optimality of $\widehat{\mathbf{w}}$ in eq.(1), we have:

$$\ell(\widehat{\mathbf{w}}) + \lambda\|\widehat{\mathbf{w}}\|_1 \leq \ell(\mathbf{w}^*) + \lambda\|\mathbf{w}^*\|_1$$

and therefore:

$$\ell(\widehat{\mathbf{w}}) - \ell(\mathbf{w}^*) \leq -\lambda\|\widehat{\mathbf{w}}\|_1 + \lambda\|\mathbf{w}^*\|_1 \quad (3)$$

By convexity of ℓ , the Cauchy-Schwarz inequality ($\forall \mathbf{a}, \mathbf{b}$) $|\langle \mathbf{a}, \mathbf{b} \rangle| \leq \|\mathbf{a}\|_1 \|\mathbf{b}\|_\infty$, and since we assume that $\lambda \geq 2\|\nabla\ell(\mathbf{w}^*)\|_\infty$, we have:

$$\begin{aligned} \ell(\widehat{\mathbf{w}}) - \ell(\mathbf{w}^*) &\geq \langle \nabla\ell(\mathbf{w}^*), \Delta \rangle \\ &\geq -\|\nabla\ell(\mathbf{w}^*)\|_\infty \|\Delta\|_1 \\ &\geq -\frac{1}{2}\lambda\|\Delta\|_1 \end{aligned} \quad (4)$$

By eq.(3) and eq.(4), it follows that $-\frac{1}{2}\lambda\|\Delta\|_1 \leq -\lambda\|\widehat{\mathbf{w}}\|_1 + \lambda\|\mathbf{w}^*\|_1$ or equivalently since $\lambda > 0$:

$$\begin{aligned} 0 &\geq -\|\Delta\|_1 + 2\|\widehat{\mathbf{w}}\|_1 - 2\|\mathbf{w}^*\|_1 \\ &\geq -\|\Delta\|_1 + 2\|\mathbf{w}^*\|_1 - 2\|\Delta_{\mathcal{K}}\|_1 + 2\|\Delta_{\mathcal{K}^c}\|_1 - 2\|\mathbf{w}^*\|_1 \\ &= -\|\Delta\|_1 - 2\|\Delta_{\mathcal{K}}\|_1 + 2\|\Delta_{\mathcal{K}^c}\|_1 \\ &= -\|\Delta_{\mathcal{K}}\|_1 - \|\Delta_{\mathcal{K}^c}\|_1 - 2\|\Delta_{\mathcal{K}}\|_1 + 2\|\Delta_{\mathcal{K}^c}\|_1 \\ &= -3\|\Delta_{\mathcal{K}}\|_1 + \|\Delta_{\mathcal{K}^c}\|_1 \end{aligned}$$

where the second line follows from eq.(2). Given the above, we have:

$$\begin{aligned}
\|\Delta\|_1 &= \|\Delta_{\mathcal{K}}\|_1 + \|\Delta_{\mathcal{K}^c}\|_1 \\
&\leq \|\Delta_{\mathcal{K}}\|_1 + 3\|\Delta_{\mathcal{K}}\|_1 \\
&= 4\|\Delta_{\mathcal{K}}\|_1 \\
&\leq 4\sqrt{k}\|\Delta_{\mathcal{K}}\|_2 \\
&\leq 4\sqrt{k}\|\Delta\|_2
\end{aligned} \tag{5}$$

which proves our claim. \square

Next, we provide the final proof.

Proof of Theorem 8.1. Let $\Delta \equiv \widehat{\mathbf{w}} - \mathbf{w}^*$. First, since we assume that $\lambda \geq 2\|\nabla\ell(\mathbf{w}^*)\|_\infty$ we can invoke Lemma 8.1 and therefore:

$$\|\Delta\|_1 \leq 4\sqrt{k}\|\Delta\|_2 \tag{6}$$

For $\mathbf{w} = \widehat{\mathbf{w}}$, by restricted strong convexity of ℓ around \mathbf{w}^* with parameters (α, τ, g) as in Definition 8.1, by eq.(6) and since $17\frac{\tau}{\alpha}kg(n, p) \leq 1$, we have:

$$\begin{aligned}
\ell(\widehat{\mathbf{w}}) - \ell(\mathbf{w}^*) - \langle \nabla\ell(\mathbf{w}^*), \Delta \rangle &\geq \alpha\|\Delta\|_2^2 - \tau g(n, p)\|\Delta\|_1^2 \\
&\geq (\alpha - 16k\tau g(n, p))\|\Delta\|_2^2 \\
&\geq (\alpha - \frac{16}{17}\alpha)\|\Delta\|_2^2 \\
&= \frac{1}{17}\alpha\|\Delta\|_2^2
\end{aligned} \tag{7}$$

By eq.(6) and eq.(7), the Cauchy-Schwarz inequality $(\forall \mathbf{a}, \mathbf{b}) \quad |\langle \mathbf{a}, \mathbf{b} \rangle| \leq \|\mathbf{a}\|_1 \|\mathbf{b}\|_\infty$, and since we assume that $\lambda \geq 2\|\nabla\ell(\mathbf{w}^*)\|_\infty$, we have:

$$\begin{aligned}
\frac{1}{17}\alpha\|\Delta\|_2^2 &\leq \ell(\widehat{\mathbf{w}}) - \ell(\mathbf{w}^*) - \langle \nabla\ell(\mathbf{w}^*), \Delta \rangle \\
&\leq -\lambda\|\widehat{\mathbf{w}}\|_1 + \lambda\|\mathbf{w}^*\|_1 + \|\nabla\ell(\mathbf{w}^*)\|_\infty\|\Delta\|_1 \\
&\leq -\lambda\|\widehat{\mathbf{w}}\|_1 + \lambda\|\mathbf{w}^*\|_1 + \frac{1}{2}\lambda\|\Delta\|_1 \\
&\leq -\lambda\|\widehat{\mathbf{w}}\|_1 + \lambda\|\widehat{\mathbf{w}}\|_1 + \lambda\|\Delta\|_1 + \frac{1}{2}\lambda\|\Delta\|_1 \\
&= \frac{3}{2}\lambda\|\Delta\|_1 \\
&\leq 6\sqrt{k}\lambda\|\Delta\|_2
\end{aligned}$$

which proves the first claim after canceling $\|\Delta\|_2$ on both sides of the inequality. The second claim can be proven by the above and eq.(6). \square

2 Application: Compressed Sensing

Assume that there is an unknown but fixed $\mathbf{w}^* \in \mathbb{R}^p$. The only way to access \mathbf{w}^* is through a *black box* that works as follows. We (somehow) generate an *input* vector $\mathbf{x}_i \in \{-1, +1\}^p$ and the black box returns an *output*:

$$y_i = \langle \mathbf{x}_i, \mathbf{w}^* \rangle + \varepsilon_i$$

where $\varepsilon_i \in \{-1, +1\}$ is a Rademacher random variable (see Definition 6.1). In the above, we know that y_i is equal to $\langle \mathbf{x}_i, \mathbf{w}^* \rangle + \varepsilon_i$, but we do not have access to \mathbf{w}^* or ε_i . We only have access to the output y_i , and of course the input \mathbf{x}_i .

The question is how many pairs (\mathbf{x}_i, y_i) are sufficient in order to recover a vector $\widehat{\mathbf{w}}$ which is *close to* \mathbf{w}^* . Assume we obtain n pairs. Let $\mathbf{X} \in \{-1, +1\}^{n \times p}$, $\mathbf{y} \in \mathbb{R}^n$ and $\boldsymbol{\epsilon} \in \{-1, +1\}^n$. Note that:

$$\mathbf{y} = \mathbf{X}\mathbf{w}^* + \boldsymbol{\epsilon} \quad (8)$$

We solve eq.(1) by using the loss function:

$$\ell(\mathbf{w}) = \frac{1}{2n} \|\mathbf{X}\mathbf{w} - \mathbf{y}\|_2^2 \quad (9)$$

For answering our question, we will have to show that the loss function ℓ fulfills the conditions of Theorem 8.1. From eq.(8) and eq.(9), we have:

$$\begin{aligned} \ell(\mathbf{w}) &= \frac{1}{2n} \|\mathbf{X}(\mathbf{w} - \mathbf{w}^*) - \boldsymbol{\epsilon}\|_2^2 \\ &= \frac{1}{2n} (\mathbf{w} - \mathbf{w}^*)^T \mathbf{X}^T \mathbf{X} (\mathbf{w} - \mathbf{w}^*) - \frac{1}{n} \boldsymbol{\epsilon}^T \mathbf{X} (\mathbf{w} - \mathbf{w}^*) + \frac{1}{2n} \boldsymbol{\epsilon}^T \boldsymbol{\epsilon} \end{aligned} \quad (10)$$

By the above, we can conclude that:

$$\ell(\mathbf{w}^*) = \frac{1}{2n} \boldsymbol{\epsilon}^T \boldsymbol{\epsilon} \quad (11)$$

$$\nabla \ell(\mathbf{w}) = \frac{1}{n} \mathbf{X}^T \mathbf{X} (\mathbf{w} - \mathbf{w}^*) - \frac{1}{n} \mathbf{X}^T \boldsymbol{\epsilon}$$

$$\nabla \ell(\mathbf{w}^*) = -\frac{1}{n} \mathbf{X}^T \boldsymbol{\epsilon} \quad (12)$$

In what follows we will assume that each entry of \mathbf{X} and $\boldsymbol{\epsilon}$ is independent and Rademacher distributed.

First Condition: $\lambda \geq 2\|\nabla \ell(\mathbf{w}^*)\|_\infty$. Assume that we set the regularization weight as follows $\lambda = 4\sqrt{\frac{\log p}{n}}$. Let $\mathbf{x}^j \in \{-1, +1\}^n$ be the j -th column of \mathbf{X} . Fix j . Note that $\frac{1}{n} \langle \mathbf{x}^j, \boldsymbol{\epsilon} \rangle = \frac{1}{n} \sum_{i=1}^n x_{ij} \varepsilon_i = \frac{1}{n} \sum_{i=1}^n z_i$ where $z_i \equiv x_{ij} \varepsilon_i$ for $i = 1 \dots n$ are independent random variables. Moreover, $z_i \in [-1, +1]$ and $\mathbb{E}_\epsilon[\frac{1}{n} \langle \mathbf{x}^j, \boldsymbol{\epsilon} \rangle \mid \mathbf{X}] = 0$ since $\mathbb{E}_\epsilon[\boldsymbol{\epsilon} \mid \mathbf{X}] = \mathbf{0}$. Thus, by Hoeffding's inequality (Corollary 2.2) and the union bound:

$$\begin{aligned} \mathbb{P}_\epsilon \left[(\exists j = 1 \dots p) \left| \frac{1}{n} \langle \mathbf{x}^j, \boldsymbol{\epsilon} \rangle \right| \geq \frac{\lambda}{2} \mid \mathbf{X} \right] &\leq 2p e^{-\frac{2n^2(\lambda/2)^2}{n2^2}} \\ &= 2p e^{-\frac{n\lambda^2}{8}} \\ &= 2p e^{-2 \log p} \\ &= 2/p \end{aligned}$$

By eq.(12) and the above, we have:

$$\begin{aligned}
\mathbb{P}_{\mathbf{X}, \epsilon} \left[\|\nabla \ell(\mathbf{w}^*)\|_\infty \geq \frac{\lambda}{2} \right] &= \mathbb{P}_{\mathbf{X}, \epsilon} \left[\left\| \frac{1}{n} \mathbf{X}^T \boldsymbol{\epsilon} \right\|_\infty \geq \frac{\lambda}{2} \right] \\
&= \mathbb{P}_{\mathbf{X}, \epsilon} \left[(\exists j = 1 \dots p) \left| \frac{1}{n} \langle \mathbf{x}^j, \boldsymbol{\epsilon} \rangle \right| \geq \frac{\lambda}{2} \right] \\
&= \mathbb{E}_{\mathbf{X}} \left[\mathbb{P}_{\epsilon} \left[(\exists j = 1 \dots p) \left| \frac{1}{n} \langle \mathbf{x}^j, \boldsymbol{\epsilon} \rangle \right| \geq \frac{\lambda}{2} \middle| \mathbf{X} \right] \right] \\
&\leq \mathbb{E}_{\mathbf{X}} [2/p] \\
&= 2/p
\end{aligned}$$

Therefore, with probability at least $1 - 2/p$ over the choice of \mathbf{X} and $\boldsymbol{\epsilon}$, we have that $\lambda \geq 2\|\nabla \ell(\mathbf{w}^*)\|_\infty$ when we use the regularization weight $\lambda = 4\sqrt{\frac{\log p}{n}}$.

Second Condition: Retricted Strong Convexity. Theorem 8.1 requires that the loss function ℓ in eq.(9) fulfills Definition 8.1. Here, we will show that indeed this is fulfilled. That is:

$$(\forall \mathbf{w} \in \mathbb{R}^p) \ell(\mathbf{w}) - \ell(\mathbf{w}^*) - \langle \nabla \ell(\mathbf{w}^*), \mathbf{w} - \mathbf{w}^* \rangle \geq \alpha \|\mathbf{w} - \mathbf{w}^*\|_2^2 - \tau g(n, p) \|\mathbf{w} - \mathbf{w}^*\|_1^2$$

Note that by eq.(10), eq.(11) and eq.(12), we have:

$$\begin{aligned}
\ell(\mathbf{w}) - \ell(\mathbf{w}^*) - \langle \nabla \ell(\mathbf{w}^*), \mathbf{w} - \mathbf{w}^* \rangle &= \frac{1}{2n} (\mathbf{w} - \mathbf{w}^*)^T \mathbf{X}^T \mathbf{X} (\mathbf{w} - \mathbf{w}^*) \\
&= \frac{1}{2n} \|\mathbf{X}(\mathbf{w} - \mathbf{w}^*)\|_2^2
\end{aligned}$$

Let $\mathbf{v} = \mathbf{w} - \mathbf{w}^*$. Our goal is to show that:

$$(\forall \mathbf{v} \in \mathbb{R}^p) \frac{1}{2n} \|\mathbf{X}\mathbf{v}\|_2^2 \geq \alpha \|\mathbf{v}\|_2^2 - \tau g(n, p) \|\mathbf{v}\|_1^2$$

Since the above is trivially fulfilled for $\mathbf{v} = \mathbf{0}$ and since if the above holds for some $\mathbf{v} \in \mathbb{R}^p$ then it also holds for $c\mathbf{v}$ for all $c \in \mathbb{R}$, we will equivalently show that:

$$(\forall \|\mathbf{v}\|_1 = 1) \frac{1}{2n} \|\mathbf{X}\mathbf{v}\|_2^2 \geq \alpha \|\mathbf{v}\|_2^2 - \tau g(n, p)$$

Fix $j \neq k$. Note that $\frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} = \frac{1}{n} \sum_{i=1}^n z_i$ where $z_i \equiv x_{ij} x_{ik}$ for $i = 1 \dots n$ are independent random variables. Moreover, $z_i \in [-1, +1]$ and we also know that $\mathbb{E}_{\mathbf{X}}[\frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik}] = 0$ since the entries of \mathbf{X} are independent and zero-mean. Thus, by Hoeffding's inequality (Corollary 2.2), the union bound and by

assuming $t = \sqrt{6 \frac{\log p}{n}}$:

$$\begin{aligned}
\mathbb{P}_{\mathbf{X}} \left[\max_{j \neq k} \left| \frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} \right| \geq t \right] &= \mathbb{P}_{\mathbf{X}} \left[(\exists j \neq k) \left| \frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} \right| \geq t \right] \\
&\leq 2 \binom{p}{2} e^{-\frac{2n^2 t^2}{n^2}} \\
&\leq p^2 e^{-\frac{nt^2}{2}} \\
&= p^2 e^{-3 \log p} \\
&= 1/p
\end{aligned}$$

Therefore, with probability at least $1 - 1/p$ over the choice of \mathbf{X} , we have that:

$$\max_{j \neq k} \left| \frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} \right| \leq \sqrt{6 \frac{\log p}{n}} \quad (13)$$

Note that since $\|\mathbf{v}\|_1 = 1$, then:

$$\begin{aligned}
\sum_{j \neq k} |v_j v_k| &\leq \sum_{j=1}^p \sum_{k=1}^p |v_j v_k| \\
&= \sum_{j=1}^p \sum_{k=1}^p |v_j| |v_k| \\
&= \|\mathbf{v}\|_1^2 \\
&= 1
\end{aligned} \quad (14)$$

Since $(\forall ij) x_{ij}^2 = 1$ and the above, we have:

$$\begin{aligned}
\frac{1}{2n} \|\mathbf{X}\mathbf{v}\|_2^2 &= \frac{1}{2n} \sum_{i=1}^n (\mathbf{X}\mathbf{v})_i^2 \\
&= \frac{1}{2n} \sum_{i=1}^n \left(\sum_{j=1}^p x_{ij} v_j \right)^2 \\
&= \frac{1}{2n} \sum_{i=1}^n \left(\sum_{j=1}^p x_{ij}^2 v_j^2 + \sum_{j \neq k} x_{ij} x_{ik} v_j v_k \right) \\
&= \frac{1}{2n} \sum_{i=1}^n \left(\sum_{j=1}^p v_j^2 + \sum_{j \neq k} x_{ij} x_{ik} v_j v_k \right) \\
&= \frac{1}{2} \|\mathbf{v}\|_2^2 + \frac{1}{2} \sum_{j \neq k} \left(\frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} \right) v_j v_k
\end{aligned}$$

$$\begin{aligned}
&\geq \frac{1}{2} \|\mathbf{v}\|_2^2 - \frac{1}{2} \max_{j \neq k} \left| \frac{1}{n} \sum_{i=1}^n x_{ij} x_{ik} \right| \sum_{j \neq k} |v_j v_k| \\
&\geq \frac{1}{2} \|\mathbf{v}\|_2^2 - \frac{1}{2} \sqrt{6 \frac{\log p}{n}}
\end{aligned}$$

where the previous-to-the-last step follows from the Cauchy-Schwarz inequality $(\forall \mathbf{a}, \mathbf{b}) |\langle \mathbf{a}, \mathbf{b} \rangle| \leq \|\mathbf{a}\|_1 \|\mathbf{b}\|_\infty$. The last step follows from eq.(13) and eq.(14).

Note that given our brief introduction at the beginning of the proof, we have shown that:

$$(\forall \mathbf{w} \in \mathbb{R}^p) \ell(\mathbf{w}) - \ell(\mathbf{w}^*) - \langle \nabla \ell(\mathbf{w}^*), \mathbf{w} - \mathbf{w}^* \rangle \geq \frac{1}{2} \|\mathbf{w} - \mathbf{w}^*\|_2^2 - \frac{1}{2} \sqrt{6 \frac{\log p}{n}} \|\mathbf{w} - \mathbf{w}^*\|_1^2$$

Therefore, we conclude that the loss function ℓ in eq.(9) fulfills Definition 8.1 with $\alpha = 1/2$, $\tau = \sqrt{6}/2$ and $g(n, p) = \sqrt{\frac{\log p}{n}}$.

Third Condition: Sufficient Number of Samples. Theorem 8.1 also requires that $17 \frac{\tau}{\alpha} k g(n, p) \leq 1$. That is:

$$17 \frac{\tau}{\alpha} k g(n, p) = 17 \sqrt{6} k \sqrt{\frac{\log p}{n}} \leq 1$$

Thus, we require $n \geq 17^2 6 k^2 \log p$.

A proof for possibly correlated Gaussian random variables can be found in [2] where they obtained results with $g(n, p) = \frac{\log p}{n}$, which is better for the required number of samples in the Third Condition.

References

- [1] S. Negahban, P. Ravikumar, M. Wainwright, and B. Yu. A unified framework for high-dimensional analysis of M-estimators with decomposable regularizers. *Statistical Science*, 2012.
- [2] G. Raskutti, M. Wainwright, and B. Yu. Restricted eigenvalue properties for correlated Gaussian designs. *Journal of Machine Learning Research*, 11(Aug):2241–2259, 2010.