Lecture 26: Left-over Hash Lemma & Bonami-Beckner Noise Operator

Suppose we have access to a sample from a probability distribution X that only has very weak randomness guarantee. For example, X is a probability distribution over the sample space {0,1}<sup>n</sup> such that H<sub>∞</sub>(X) ≥ k. That is, the output of X is very unpredictable and for all x ∈ {0,1}<sup>n</sup>

$$\mathbb{P}\left[\mathbb{X}=x\right] \leqslant \frac{1}{2^k} = \frac{1}{K}$$

 Our objective is to generate uniform random bits from any distribution with H<sub>∞</sub>(X) ≥ k

- Ideally, we will prefer to have <u>one</u> function
   f: {0,1}<sup>n</sup> → {0,1}<sup>m</sup> such that it can its output f(X) is close
   to the uniform distribution U<sub>m</sub> (the uniform distribution over
   {0,1}<sup>m</sup>)
- However, we shall show that it is impossible that <u>one function</u> can extract random bits from <u>all</u> high min-entropy sources. This impossibility is in the strongest possible sense.
- We shall show that for every extraction function
   f: {0,1}<sup>n</sup> → {0,1}, there exists a min-entropy source X such
   that H<sub>∞</sub>(X) ≥ n - 1 such that f(X) is constant. We cannot
   even extract one random bit from sources with (n - 1)
   min-entropy.

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The proof is as follows. Consider S<sub>0</sub> = f<sup>-1</sup>(0) and S<sub>1</sub> = f<sup>-1</sup>(1). Note that either S<sub>0</sub> or S<sub>1</sub> has at least 2<sup>n-1</sup> entries. Suppose without loss of generality, |S<sub>0</sub>| ≥ 2<sup>n-1</sup>. Consider X, the uniform distribution over the set S<sub>0</sub>. Note that P[X = x] ≤ 1/(2<sup>n-1</sup>). We have H<sub>∞</sub>(X) ≥ n - 1.

## Definition (Universal Hash Function Family)

Let  $\mathcal{H} = \{h_1, h_2, \ldots, h_\alpha\}$  be a collection of hash functions such that, for each  $1 \leq i \leq \alpha$ , we have  $h_i \colon \{0,1\}^n \to \{0,1\}^m$ . Let  $\mathbb{H}$  be a probability distribution over the hash functions in  $\mathcal{H}$ . The family  $\mathcal{H}$  is a *universal hash function family* with respect to the probability distribution  $\mathbb{H}$  if it satisfies the following condition. For all distinct inputs  $x, x' \in \{0,1\}^n$ , we have

$$\mathbb{P}\left[h(x)=h(x')\colon h\sim\mathbb{H}
ight]\leqslantrac{1}{2^m}=rac{1}{M}$$

- Recall that we have seen that it is impossible for a deterministic function to extract even one random bit from sources with (n-1) bits of min-entropy.
- We shall now show that choosing a hash function from a universal hash function family suffices

## Theorem (Left-over Hash Lemma)

Let  $\mathcal{H}$  be a universal hash function family  $\{0,1\}^n \to \{0,1\}^m$  with respect to the probability distribution  $\mathbb{H}$  over  $\mathcal{H}$ . Let  $\mathbb{X}$  be any min-entropy source over  $\{0,1\}^n$  such that  $H_{\infty}(\mathbb{X}) \ge k$ . Then, we have

$$\mathrm{SD}\left((\mathbb{H}(\mathbb{X}),\mathbb{H}),(\mathbb{U}_m,\mathbb{H})
ight)\leqslant rac{1}{2}\sqrt{rac{M}{K}}$$

Remark. Note that we are claiming that 𝔑(𝔅) is close to the uniform distribution 𝔅<sub>m</sub> over {0,1}<sup>m</sup> even given the hash function 𝔅.

• The proof proceeds in the following steps.

$$2SD ((\mathbb{H}(\mathbb{X}), \mathbb{H}), (\mathbb{U}_m, \mathbb{H}))$$

$$=\mathbb{E} \left[2SD ((\mathbb{H}(\mathbb{X})|\mathbb{H} = h), (\mathbb{U}_m|\mathbb{H} = h)) : h \sim \mathbb{H}\right]$$

$$=\mathbb{E} \left[2SD (h(\mathbb{X}), \mathbb{U}_m) : h \sim \mathbb{H}\right]$$

$$\leq \mathbb{E} \left[\ell_2 \left(\operatorname{Bias}_{h(\mathbb{X})} - \operatorname{Bias}_{\mathbb{U}_m}\right) : h \sim \mathbb{H}\right]$$

$$=\mathbb{E} \left[\sqrt{\sum_{S \in \{0,1\}^m} \operatorname{Bias}_{h(\mathbb{X})}(S)^2 - 1} : h \sim \mathbb{H}\right]$$

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The last inequality is due to Jensen's inequality.

• Let us continue our simplification.

$$2 \operatorname{SD} \left( (\mathbb{H}(\mathbb{X}), \mathbb{H}), (\mathbb{U}_m, \mathbb{H}) \right)$$
$$\leq \sqrt{\mathbb{E} \left[ \sum_{S \in \{0,1\}^m} \operatorname{Bias}_{h(\mathbb{X})}(S)^2 - 1 \colon h \sim \mathbb{H} \right]}$$
$$= \sqrt{\mathbb{E} \left[ \sum_{S \in \{0,1\}^m} \operatorname{Bias}_{h(\mathbb{X})}(S)^2 \colon h \sim \mathbb{H} \right] - 1}$$
$$= \sqrt{\mathbb{E} \left[ M \cdot \operatorname{Col} \left( h(\mathbb{X}), h(\mathbb{X}) \right) \colon h \sim \mathbb{H} \right] - 1}$$

- Note that one sample of h(X) collides with a second sample of h(X) due to the following cases
  - The first sample of X collides with the second sample of X. Since, H<sub>∞</sub>(X) ≥ k, we have

$$\operatorname{Col}(\mathbb{X},\mathbb{X})\leqslant rac{1}{K}$$

② If the first and the second samples from X are different, then they collide with probability  $≤ \frac{1}{M}$  when  $h ∼ \mathbb{H}$ .

Overall, by union bound, we get that

$$\mathbb{E}\left[\operatorname{Col}\left(h(\mathbb{X}),h(\mathbb{X})
ight):h\sim\mathbb{H}
ight]\leqslantrac{1}{K}+rac{1}{M}$$

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• Substituting this estimation, we obtain

$$\begin{split} &2\mathrm{SD}\left((\mathbb{H}(\mathbb{X}),\mathbb{H}),(\mathbb{U}_m,\mathbb{H})\right) \\ \leqslant &\sqrt{\mathbb{E}\left[M\cdot\mathrm{Col}\left(h(\mathbb{X}),h(\mathbb{X})\right):h\sim\mathbb{H}\right]-1} \\ =&\sqrt{M\cdot\left(\frac{1}{K}+\frac{1}{M}\right)-1} = \sqrt{\frac{M}{K}} \end{split}$$

• Note that this result says that we must ensure *m* < *k* for the output of the extraction to be close to the uniform distribution

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- Today, we shall introduce the basics of the "noise operator"
- This operator is crucial to one of the most powerful technical tools in Fourier Analysis, namely, the Hypercontractivity

• Let  $\mathbb{N}_{\varepsilon}$  be a probability distribution over the sample space  $\{0,1\}^n$  such that

$$\mathbb{P}\left[\mathbb{N}_{\varepsilon}=x\right]=(1-\varepsilon)^{n-|x|}\varepsilon^{|x|}$$

Here |x| represents the number of 1s in x (or, equivalently, the Hamming weight of x)

- Intuitively, imagine a channel through which 0<sup>n</sup> is fed as input. The channel converts each bit independently as follows. It converts 0 → 1 with probability ε; and 1 → 0 with probability (1 - ε). Note that the probability of the output being x is (1 - ε)<sup>n-|x|</sup>ε<sup>|x|</sup>
- Our objective is to prove that

$$\operatorname{Bias}_{\mathbb{N}_{\varepsilon}}(S) = (1 - 2\varepsilon)^{|S|}$$

We shall prove this result using a highly modular and elegant approach

 For 1 ≤ i ≤ n, let N<sub>ε,i</sub> be the probability distribution defined below

$$\mathbb{P}\left[\mathbb{N}_{arepsilon,i}=x
ight]=egin{cases} (1-arepsilon), & ext{if } x=0^n\ arepsilon, & ext{if } x=\delta_i\ 0, & ext{otherwise} \end{cases}$$

Intuitively, 0<sup>n</sup> is fed through a channel. All bits except the *i*-th bit are left unchanged. The *i*-th bit is converted as follows. It maps 0 → 1 with probability ε; and 0 → 0 with probability (1 - ε).

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## Computation of the Bias

• Let us compute the bias of this distribution. For any  $S \in \{0,1\}^n$ , note that, if  $S_i = 0$ , we have

$$\operatorname{Bias}_{\mathbb{N}_{\varepsilon,i}}(S) = 1$$

For any  $S \in \{0,1\}$ , if  $S_i = 1$ , we have

$$\operatorname{Bias}_{\mathbb{N}_{\varepsilon,i}}(S) = (1-\varepsilon) - \varepsilon = (1-2\varepsilon)$$

• Succinctly, we can express this as

$$\operatorname{Bias}_{\mathbb{N}_{\varepsilon,i}}(S) = (1-2\varepsilon)^{S_i}$$

• So, we can conclude that

$$\operatorname{Bias}_{\bigoplus_{i=1}^n \mathbb{N}_{arepsilon,i}}(S) = (1-2arepsilon)^{\sum_{i=1}^n S_i} = (1-2arepsilon)^{|S|}$$

• It is left as an exercise to prove that the distribution  $\mathbb{N}_{\varepsilon}$  is identical to the distribution  $\bigoplus_{i=1}^{n} \mathbb{N}_{\varepsilon,i}$ 

## Noisy Version of a Function

- Let  $f: \{0,1\}^n \to \mathbb{R}$  be any function
- Define the noisy version of f as follows

$$\widetilde{f}(x) = T_
ho(x) \mathrel{\mathop:}= \mathbb{E}\left[f(x+e) \colon e \sim \mathbb{N}_arepsilon
ight],$$

where  $\rho = 1-2\varepsilon$ 

• So, we have

$$\widetilde{f}(x) = \sum_{e \in \{0,1\}^n} \mathbb{N}_{\varepsilon}(e) f(x+e) = N(\mathbb{N}_{\varepsilon} * f)$$

Equivalently, we have  $\tilde{f} = \mathbb{N}_{\varepsilon} \oplus f$  (we emphasize that f need not be a probability distribution to use the notation of  $\oplus$  of two functions)

Therefore, we get

$$\operatorname{Bias}_{\widetilde{f}}(S) = \operatorname{Bias}_{\mathbb{N}_{\varepsilon}}(S) \cdot \operatorname{Bias}_{f}(S) = \rho^{|S|} \operatorname{Bias}_{f}(S)$$

• That is, we conclude that

$$\widehat{T_{\rho}(f)}(S) = \rho^{|S|} \widehat{f}(S)$$