Lecture 03: Balls & Bins (Birthday Paradox, Max-Load)

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- There are *m* balls and *n* bins
- Each balls is independently thrown into a bin that is chosen uniformly at random
- $(\mathbb{T}_1, \dots, \mathbb{T}_m)$ be the joint distribution such that \mathbb{T}_i represents the bin into which the *i*-th ball is thrown

- For 1 ≤ i < j ≤ m, let X_{i,j} be the indicator variable for the event that the *i*-th ball and the *j*-th ball fall into the same variable, i.e., indicator variable for the event T_i = T_j
- We are interested in computing the "expectation of the random variable $\mathbb{X}_{i,j}$ "

$$\mathbb{E}\left[\mathbb{X}_{i,j}\right] = \mathbb{P}\left[\mathbb{X}_{i,j} = 1\right] = \mathbb{P}\left[\mathbb{T}_i = \mathbb{T}_j\right]$$
$$= \sum_{t=1}^n \mathbb{P}\left[\mathbb{T}_i = \mathbb{T}_j = t\right] = n \cdot 1/n^2 = 1/n$$

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Expected Number of Collisions

• Let
$$\mathbb{X} = \sum_{1 \leqslant i < j \leqslant m} \mathbb{X}_{i,j}$$

- $\bullet\,$ The random variable $\mathbb X$ counts the number of collisions that occur
- We are interested in the expected number of collisions

$$\mathbb{E}\left[\mathbb{X}\right] = \mathbb{E}\left[\sum_{1 \leq i < j \leq m} \mathbb{X}_{i,j}\right]$$
$$= \sum_{1 \leq i < j \leq m} \mathbb{E}\left[\mathbb{X}_{i,j}\right] \quad \text{By the Linearity of Expectation}$$
$$= \binom{m}{2} \frac{1}{n}$$

- Note that for $m=\sqrt{2n},$ we have $\mathbb{E}\left[\mathbb{X}
 ight]=1$
- The "average number of collisions" at $m = \sqrt{2n}$ is 1, but how does the probability of this event behave like?

There are m people in a room. Assume that the birthday of people are distributed uniformly at random over the 365 days in the year. What is the number of people m that ensures that two people share a birthday with probability 0.9?

- We want to find the value of m such that throwing m balls in n = 365 bins ensures a collision with probability 0.9
- Let NoColl_{\leqslant t} represent the probability that $\mathbb{T}_1,\ldots,\mathbb{T}_t$ are all distinct
- Note that

$$\mathbb{P}\left[\mathsf{NoColl}_{\leqslant t}\right] = 1 \cdot \left(1 - \frac{1}{n}\right) \cdot \left(1 - \frac{2}{n}\right) \cdots \left(1 - \frac{t - 1}{n}\right)$$
$$= \prod_{t=0}^{t-1} \left(1 - \frac{t}{n}\right)$$

Birthday Paradox – Balls and Bins Versions

- We are interested in finding *m* such that the probability of collision is high
- Alternately, we are interested in showing the probability of NoColl_{≤m} is small

$$\mathbb{P}\left[\mathsf{NoColl}_{\leq m}\right] = \prod_{t=1}^{m-1} \left(1 - \frac{t}{n}\right)$$
$$\leq \prod_{t=1}^{m-1} \exp\left(-\frac{t}{n}\right) = \exp\left(-\sum_{t=0}^{m-1} \frac{t}{n}\right)$$
$$= \exp\left(-(m-1)\frac{m}{2n}\right)$$

• Substituting $m = c\sqrt{n}$, for a suitable constant c > 0, ensures that $\exp\left(-(m-1)m/2n\right) \leqslant 0.1$

Birthday Paradox – Balls and Bins Versions

- We are interested in finding out *m* such that we can throw *m* balls without getting a collision, with high probability
- That is, we are interested in showing that the probability of NoColl_{≤m} is high

$$\mathbb{P}\left[\operatorname{NoColl}_{\leq m}\right] = \prod_{t=1}^{m-1} \left(1 - \frac{t}{n}\right) \ge \prod_{t=0}^{m-1} \exp\left(-\frac{t}{n} - \frac{t^2}{n^2}\right)$$
$$= \exp\left(-\sum_{t=0}^{m-1} \frac{t}{n} - \sum_{t=0}^{m-1} \frac{t^2}{n^2}\right)$$
$$= \exp\left(-\frac{m(m-1)}{2n} - \frac{m(m-0.5)(m-1)}{3n^2}\right)$$

• For $m = d\sqrt{n}$, the first term in the exponent dominates and the second term is o(1)

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• For a constant d > 0 we can ensure that the final probability term is ≥ 0.9 s

Conclusion: As *m* increases from $d\sqrt{n}$ to $c\sqrt{n}$ the probability of no-collisions transitions from 0.9 to 0.1.

Recommended: Plot the probability of no-collisions for m = 1 to m = n, for large values of n. How quickly does the probability transition from "high" to "low" as n increases?

Number of Empty Bins

- Let X_i represent the indicator variable for the *i*-th bin being empty, i.e., the indicator of the event: T_j ≠ i, for all i ∈ {1,..., m}
- Note that

$$\mathbb{E}\left[\mathbb{X}_i\right] = \mathbb{P}\left[\mathbb{X}_i = 1\right] = \left(1 - \frac{1}{n}\right)^n$$

• Let $\mathbb{X} = \sum_{t=0}^{n} \mathbb{X}_{i}$ represent the number of empty bins

$$\mathbb{E}\left[\mathbb{X}\right] = \mathbb{E}\left[\sum_{t=1}^{n} \mathbb{X}_{i}\right] = \sum_{t=1}^{n} \mathbb{E}\left[\mathbb{X}_{i}\right] = n\left(1 - \frac{1}{n}\right)^{m} \approx n \exp\left(-\frac{m}{n}\right)$$

For m = n, we expect (roughly) n/e empty bins! For m = n log n, we expect (roughly) 1 empty bin.

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Probability of a Bin containing k Balls

- There are $\binom{m}{k}$ ways of choose the balls (indexed by) $1 \leq i_1 < i_2 < \cdots < i_k \leq m$ that fall in the bin
- The probability that these balls fall into the bin is $\frac{1}{n^k}$
- The probability that other balls fall outside is $\left(1-rac{1}{n}
 ight)^{m-k}$
- Let $X_{i,=k}$ represent the indicator variable that bin *i* contains exactly *k* balls
- Note that, we have

$$\mathbb{P}\left[\mathbb{X}_{i,=k}\right] = \binom{m}{k} \frac{1}{n^k} \left(1 - \frac{1}{n}\right)^{m-k} \tag{1}$$

Max Load

- Let L_i represent the load of the *i*-th bin, i.e., the number of balls in the *i*-th bin
- Note that \mathbb{L}_i is the random variable $|\{k \colon \mathbb{T}_k = i\}|$
- $\bullet\,$ Let $\mathbb M$ be the maximum load of the bins
- That is, \mathbb{M} is the random variable max $\{\mathbb{L}_1, \dots, \mathbb{L}_n\}$
- ullet We are interested in understanding how $\mathbb{E}\left[\mathbb{M}
 ight]$ behaves like

Theorem (Max Load)

For m = n, we have

$$\mathbb{E}\left[\mathbb{M}\right] = \Theta\left(\frac{\log n}{\log\log n}\right)$$

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- First, we want to show that \mathbb{M} is $\leqslant c \frac{\log n}{\log \log n}$ with ≈ 1 probability
- This will imply that $\mathbb{E}[\mathbb{M}]$ is upper bounded by (roughly) $c \frac{\log n}{\log \log n}$
- This is known as the "First Moment Technique"

Analysis of Max Load (Upper Bound)

Let X_{i,≥k} be the indicator variable for bin i getting ≥ k balls
Note that

$$\mathbb{P}\left[\mathbb{X}_{i, \geq k} = 1\right] \leqslant \binom{m}{k} \frac{1}{n^k}$$

- Think: Why is this true?
- We will upper bound this probability further

$$\mathbb{P}\left[\mathbb{X}_{i,\geq k}=1\right] \leqslant \binom{m}{k} \frac{1}{n^k} \leqslant \left(\frac{m}{n}\right)^k \frac{1}{k!}$$

• By union bound:

$$\mathbb{P}\left[\exists i\in[n]\colon\mathbb{X}_{i,\geqslant k}=1
ight]\leqslant\left(rac{m}{n}
ight)^krac{n}{k!}$$

Balls & Bins

Analysis of Max Load (Upper Bound)

$$\mathbb{P}\left[\exists i \in [n] \colon \mathbb{X}_{i, \geq k^*} = 1\right] \leqslant \frac{1}{n}$$

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Negating, we get:

$$\mathbb{P}\left[\forall i \in [n] \colon \mathbb{X}_{i, \geq k^*} = 0\right] \ge 1 - \frac{1}{n}$$

- Note that the event " $\forall i \in [n]$: $\mathbb{X}_{i, \geq k^*} = 0$ " implies the event " $\mathbb{M} < k^*$ "
- So, we have $\mathbb{P}\left[\mathbb{M} < k^*\right] \ge 1 \frac{1}{n}$
- This implies that $\mathbb{E}\left[\mathbb{M}\right] \leqslant \left(1 \frac{1}{n}\right)(k^* 1) + \frac{1}{n} \cdot n \leqslant k^*$

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We are interested in showing that

$$\mathbb{E}\left[\mathbb{M}\right] \geqslant d \frac{\log n}{\log \log n}$$

• There are multiple ways to show this. In particular, we can use a "Second Moment Technique" to prove this result. One of the reading materials proves the result using this technique. We will, instead, use a more general technique that shows a close connection between the balls-and-bins problem and its approximation using independent Poisson Distributions

Rough Probability Calculation

Recall Equation 1. The probability of a bin to have k balls is

$$\mathbb{P}\left[\mathbb{X}_{i,=k}\right] = \binom{m}{k} \frac{1}{n^{k}} \left(1 - \frac{1}{n}\right)^{m-k}$$
$$= \binom{m}{k} \left(\frac{1}{n\left(1 - \frac{1}{n}\right)}\right)^{k} \left(1 - \frac{1}{n}\right)^{m}$$
$$\approx \frac{1}{k!} \left(\frac{m}{n-1}\right)^{k} \left(1 - \frac{1}{n}\right)^{m}$$
$$\approx \exp\left(-m/n\right) \cdot \frac{\left(m/n\right)^{k}}{k!}$$

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• Let $\mathbb {Y}$ be the distribution over the sample space $\{0,1,2,\dots\}$ such that

$$\mathbb{P}\left[\mathbb{Y}=k\right] = \exp(-\mu)\frac{\mu^k}{k!}$$

- $\bullet\,$ Prove: This is a probability distribution with mean μ
- ullet This distribution is the Poisson Distribution with mean μ

Intuition of the Approximation

- Reality: The load distribution of n bins when m balls are thrown is represented by the joint random variables (L₁,..., L_n)
- Approximation: Consider the distribution $(\mathbb{Y}^{(1)}, \ldots, \mathbb{Y}^{(n)})$, where each $\mathbb{Y}^{(i)}$ is an independent Poisson distribution with mean m/n

Theorem (Intuitive: Poisson Approximation)

If f is a "well-behaved" function, then

$$\mathbb{E}\left[f(\mathbb{L}_1,\ldots,\mathbb{L}_n)\right] \lesssim \mathbb{E}\left[f(\mathbb{Y}^{(1)},\ldots,\mathbb{Y}^{(n)})\right]$$

Example: We want to show that the max-load is $\geq d \frac{\log n}{\log \log n}$ with high probability. So, we choose f as the indicator variable for the event that the maximum of the inputs is $< d \frac{\log n}{\log \log n}$. Then we show that the $\mathbb{E}\left[f(\mathbb{Y}^{(1)}, \ldots, \mathbb{Y}^{(n)})\right]$ is "small." So, we have $\mathbb{E}\left[f(\mathbb{L}_1, \ldots, \mathbb{L}_n)\right]$ is also "small."

Think:

- How many balls need to be thrown so that every bin has at least one ball?
- How many balls need to be thrown so that every bin has at least *r* balls?

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