# A Lower-Variance Randomized Algorithm for Approximate String Matching* 

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#### Abstract

Several randomized algorithms make use of convolution to estimate the score vector of matches between a text string of length $N$ and a pattern string of length $M$, i.e., the vector obtained when the pattern is slid along the text, and the number of matches is counted for each position. These algorithms run in deterministic time $O(k N \log M)$, and find an unbiased estimator of the scores whose variance is $(M-c)^{2} / k$ where $c$ is the actual score; here $k$ is an adjustable parameter that provides a tradeoff between computation time and lower variance. This paper provides an algorithm that also runs in in deterministic time $O(k N \log M)$ but achieves a lower variance of $\min (M / k, M-c)(M-c) / k$. For all score values $c$ that are less than $M-(M / k)$, our variance is essentially a factor of $k$ smaller than in previous work, and for $M-(M / k)<c \leq M$ it matches the previous work's variance. As in the previous work, our estimator is unbiased, and we make no assumption about the probabilistic characteristics of the input, or about the size of the alphabet, and our solution extends to string matching with classes, class complements, "never match" and "always match" symbols, to the weighted case and to higher dimensions.


[^0]Keywords: convolution, FFT, approximate string matching, randomized algorithms.

## 1 Introduction

Given a text string $T=t_{0} t_{1} \ldots t_{N-1}$ and a pattern string $P=p_{0} p_{1} \ldots p_{M-1}$ over an alphabet $\Sigma$ of size $\sigma$, we address the problem of computing the score vector of matches between $T$ and $P$. This is defined as the vector $C$ whose $i$ th component $c_{i}$ is the number of matches between the text and the pattern when the first letter of the pattern is positioned in front of the $i$ th letter of the string.

We avoid repeating the discussion of the applicability of the above problems, which is amply discussed in the cited references and elsewhere in the literature.

As in $[2,10,4]$, we seek a randomized algorithm of Monte-Carlo type to compute an unbiased estimate of the score vector, and as in the previous work we make crucial use of convolution, specifically, of the well-known fact that by doing $N / M$ convolutions one can compute all the $N-M$ scalar products between a vector of length $M$ and all contiguous length- $M$ subvectors of another vector of length $N$ (see ,e.g., $[1,9]$ ).

Also as in the previous work, the algorithm we propose here proceeds by computing and averaging $k$ independent equally distributed estimates for the score vector, and has the property that the expected value of the averaged estimator is equal to the exact value: The expected value of the $i$ th component $\widehat{c_{i}}$ of the estimate $\widehat{C}$ of the score vector equals $c_{i}$. Our improvement is that the variance is now $\min (M / k, M-c)(M-c) / k$, whereas it was $\left(M-c_{i}\right)^{2} / k$ in the previous work, $k$ being a parameter that can be tuned to achieve a desired level of accuracy. The time complexity is the same as in the previous work: $O(k N \gamma(M) / M)$, where $\gamma(M)$ is the time needed to perform the convolution of two vectors of length $M$. As is commonly done, we replace $\gamma(M)$ by $M \log M$, which corresponds to the computational model where an arithmetic operation takes constant time. This makes the time complexity $O(k N \log M)$. The following theorem summarizes our main result.

Theorem 1. An estimate for the score $C$ between a text string of length $N$ and a pattern string of length $M$ can be computed by a Monte-Carlo algorithm in time $O(k N \log M)$, where $k$ is the number of iterations in the algorithm. The randomized result has mean $C$ and each entry has a variance bounded by $\min (M / k, M-c)(M-c) / k$.

## 2 Previous Work and New Contribution

The problem of computing the score vectors has a long history, in which the goal has always been to avoid the naive brute-force $O(M N)$ solution. Fischer and Paterson used convolution to solve the special case of finding all exact occurrences (i.e., scores that equal exactly $M$ ) in the presence of "always match" symbols [7], for the case of constant-size alphabet. Abrahamson and Kosaraju independently extended the algorithm by Fischer and Paterson into a deterministic algorithm for computing the vector $C$ in time complexity of $O(\min (|\Sigma|, \sqrt{M / \log M}) N \log M)[1,9]$ for any alphabet size. An algorithm of Baeza-Yates and Gonnet, that works remarkably well in practice, solves the problem in time $O(N M \log M / \log N)$ [5], which is better than $O(N \log M)$ for very small $M$, i.e., $M=o(\log N)$. Another algorithm, by Baeza-Yates and Perleberg, solves the problem in time $O\left(N M f_{\max }\right)$ where $f_{\max }$ is the maximal occurrence frequency of symbols in the pattern [6].

An algorithm of deterministic time $O(N \log M)$ was given in [3], whose analysis depends on some restrictive assumptions on the probabilistic characteristics of the input, namely the Bernoulli model. Karloff studied the case when the alphabet is not known beforehand and gave a deterministic $O\left(N \log ^{3} M\right)$ time algorithm for estimating all the scores of mismatches [8], and also provided a randomized variant of deterministic time complexity $O\left(N \log ^{2} M\right)$. Karloff's estimator is intentionally biased in order to guarantee not to overestimate the number of mismatches by more than a constant multiplicative factor. The method apparently cannot be modified to estimate the number of matches (rather than of mismatches).

The work that is closest to this paper is in $[2,10,4]$. In these as in this paper, the algorithm runs in deterministic time $O(k N \log M)$. The main difference is that in the present paper the variance goes down by a factor of essentially $k$, for all values of a score that are $\leq M-(M / k)$ (for score values that are greater the variance bound is the same as in the previous work).

The algorithm in [2] was based on the idea of mapping the alphabet symbols into the primitive $\sigma$ th roots of unity, and then using convolution, relying on the nullity of the sum of all the $\sigma$ th roots of unity for its unbiastedestimator property. The process was repeated $k$ times, each with a different random mapping, and the average of the $k$ results was taken to be the answer. In the follow-on work in $[10,4]$ different mappings were used. The scheme in [10] is particularly well suited for small alphabet sizes $(\leq 20)$.

The main new idea here is to do a judicious pre-processing step that also takes $O(k N \log M)$ time, and then to prove that it results in the claimed im-
proved variance. That is the main new ingredient, and in fact What follows the pre-processing step could be any of the schemes in $[2,10,4]$, although here we chose a different variant for simplicity (of both the presentation and the variance proof). The next sections present the algorithm and analyze its properties.

## 3 The Algorithm

The steps of the algorithm are given below, where we assume that the symbols in $T$ that do not occur in $P$ have been replaced by a special "match nothing" symbol \#; this means that the reduced alphabet (which we still denote $\Sigma$ ) has a size $\sigma$ that is at most the length of $M$ plus one.

## Score-Estimation Algorithm:

INPUT: a text $T=t_{0} \ldots t_{N-1}$; a pattern $P=p_{0} \ldots p_{M-1}$ where the $t_{i}$ 's and the $p_{i}$ 's are letters from $\Sigma$; an integer $k$ that determines the desired tradeoff between time and accuracy.

OUTPUT: an estimate for the score vector $C$.

1. Compute in time $O(M+\sigma \log \sigma)$ the number of occurrences in $P$ of every symbol of $\Sigma$, and let $\Sigma^{\prime}$ denote the subset of $\Sigma$ whose symbols occur more than $M / k$ times in $P$. Note that $\left|\Sigma^{\prime}\right| \leq k$.
2. For every symbol $s$ in $\Sigma^{\prime}$, compute the exact contribution of symbol $s$ to the score vector $C$. It is well known that this can done with $N / M$ convolutions for each symbol, hence in $O(N \log M)$ time per symbol (see, e.g., $[1,9]$ ). As this is done for all the symbols of $\Sigma^{\prime}$, it takes a total of $O\left(\left|\Sigma^{\prime}\right| N \log M\right)$ time. The rest of the algorithm deals with estimating the contribution of the symbols in $\Sigma-\Sigma^{\prime}$.

Note. As stated earlier, the rest of the agorithm could have used any of the schemes in $[2,10,4]$. The variant we chose is convenient for the proof of the lower variance.
3. For $\ell=1,2, \cdots, k$ :
(a) From the text $T$ and pattern $P$, obtain vectors $T_{\ell}$ and (respectively) $P_{\ell}$ as follows. First, replace in $T$ and $P$ every occurrence of a symbol from $\Sigma^{\prime} \cup\{\#\}$ with a 0 . Then, for each symbol $s \in \Sigma-\Sigma^{\prime}-\{\#\}$, flip a fair coin and, if its outcome is head (tail), replace in $T$ and $P$ every occurrence of $s$ with a +1 (resp., -1).
(b) Compute the vector $C_{\ell}$ such that, for $0 \leq i \leq N-M-1, C_{\ell}[i]$ is the scalar product of $P_{\ell}$ with the sub-vector of $T_{\ell}$ of length $M$ that starts at position $i$ in $T_{\ell}$; in other words:

$$
C_{\ell}[i]=\sum_{j=0}^{M-1} P_{\ell}[j] * T_{\ell}[i+j]
$$

Implementation note: This can be done using $N / M$ convolutions, where the $i$ th such convolution is between the following two vectors of size $2 M$ each: (i) the reverse of $P_{\ell}$ padded with $M$ zeroes; and (ii) the sub-vector of $T_{\ell}$ of length $2 M$ that starts at position $i M$ in $T_{\ell}$.
4. Compute the vector $\widehat{C}=\sum_{\ell=1}^{k} C^{\ell} / k$ as the estimate of the contribution to $C$ of symbols in $\Sigma-\Sigma^{\prime}$. Adding to this the contribution of symbols in $\Sigma^{\prime}$ as computed in Step 2 gives the desired estimate of $C$.

## 4 Analysis of the Output Estimate

That the estimator of a $c_{i}$ for that particular iteration is unbiased is obvious, because each one of the $M$ terms of the summation that it makes up has zero mean if the two symbols are different and 1 if they are same. So we henceforth focus on the variance.

We first analyze the variance for one of the $k$ iterations, and then divide by $k$ to obtain the variance for their average. The variance is the sum of the $M^{2}$ entries of the $M \times M$ covariance matrix for the $M$ terms of the summation that makes up a this iteration's estimate of $c_{i}$. For a term that corresponds to equal symbols, the corresponding row and column of the covariance matrix are zero, and because there are $c_{i}$ such terms the submatrix that contributes to nonzero entries is no larger than $\left(M-c_{i}\right) \times\left(M-c_{i}\right)$ (it can be smaller because of the symbols from $\Sigma^{\prime}$, whose involvement in any of the terms causes the value of that term to be zero, and its row and column to also be all zeroes in the covariance matrix). A nonzero entry of the covariance sub-matrix contains a 1 , and has a row (resp., column) that is a term whose corresponding 2 symbols are unequal and are both from $\Sigma-\Sigma^{\prime}$. For any such term (say, arising out of the alignment of distinct symbols $a$ and $b$ ), its covariance with any other such term is zero unless the other term also involves the same two symbols $a$ and $b$. But the number of
such repetitions of symbols $a$ and $b$ is bounded by $M / k$ (by its definition $\Sigma-\Sigma^{\prime}$ excludes symbols that occur more than $M / k$ times in $\left.P\right)$. This means that no such term can have more than $M / k$ nonzero covariances with other terms, whence the upper bound of $\left(M-c_{i}\right)(M / k)$ on the number of nonzero entries of the covariance matrix. Of course $\left(M-c_{i}\right)^{2}$ is also an upper bound, and combining these two gives us the desired $\min \left(M / k, M-c_{i}\right)\left(M-c_{i}\right)$ bound on the variance of the estimator of $c_{i}$ for that iteration. Therefore the variance of the overall estimator (the average over the $k$ iterations) is $\min \left(M / k, M-c_{i}\right)\left(M-c_{i}\right) / k$.

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