

Properties of a Poisson process

1. The expected number of arrivals in an interval of t steps is λt .
2. Sum of Poisson processes is a Poisson process with sum of rates.
3. If a Poisson process is split randomly, the two processes are Poisson.

Exponential Distribution

$$X \sim \text{Exp}(\mu) \Rightarrow \Pr(X \leq s) = 1 - e^{-\mu s}$$

$$E[X] = 1/\mu$$

The exponential distribution is **memoryless**:

$$\begin{aligned} \Pr(X > t + \tau | X > t) &= \frac{\Pr(X > t + \tau)}{\Pr(X > t)} \\ &= \frac{e^{-\mu(t+\tau)}}{e^{-\mu t}} = e^{-\mu\tau} = \Pr(X > \tau) \end{aligned}$$

Gamma Distribution

A Gamma distribution with parameters (α, β) has a density function given by:

$$f(x) = \frac{\beta^{-\alpha} x^{\alpha-1} e^{-x/\beta}}{\Gamma(\alpha)} \text{ if } x > 0$$

= 0 otherwise

where $\Gamma(\alpha)$ is the gamma function defined by

$$\Gamma(z) = \int_0^{\infty} t^{z-1} e^{-t} dt \text{ for any real number } z > 0.$$

Some properties of the gamma function are:

1. $\Gamma(z + 1) = z\Gamma(z)$ for any $z > 0$.
2. $\Gamma(k + 1) = k!$ for any nonnegative integer k .

Properties of the Gamma Distribution

1. Mean of the Gamma distribution is $\alpha\beta$ and the variance is $\alpha\beta^2$.
2. The $Exp(\beta)$ and $\Gamma(1, 1/\beta)$ are the same.
3. If X_1, \dots, X_m are independent $Exp(\beta)$ r.v.s, then $\sum_{i=1}^m X_i$ is $\Gamma(m, 1/\beta)$ – also called the *m-Erlang* distribution.

Interarrival distributions

Let X_n ($n \geq 1$) denote the time from the $(n-1)$ th to the n th arrival.

The sequence $\{X_n, n = 1, 2, \dots\}$ is the **sequence of interarrival times**.

$$\Pr(X_1 > t) = \Pr(N(t) = 0) = e^{-\lambda t}$$

Hence X_1 has exponential distribution with mean $1/\lambda$.

$$\Pr(X_2 > t) = E[\Pr(X_2 > t | X_1)]$$

However,

$$\begin{aligned}\Pr(X_2 > t | X_1 = s) &= \Pr(0 \text{ events in } (s, s+t] | X_1 = s) \\ &= e^{-\lambda t}\end{aligned}$$

If the arrival process is Poisson with rate λ , then X_n , $n = 1, 2, \dots$ are independently and identically distributed exponential random variable with mean $1/\lambda$.

Waiting Time Distribution

S_n : arrival time of the n th event - called the *waiting time* of the n th event.

$$S_n = \sum_{i=1}^n X_i \quad n \geq 1$$

S_n has a *Gamma* distribution with parameters n and λ .

$$N(t) \geq n \equiv S_n \leq t$$

Hence,

$$F_{S_n}(t) = \sum_{j=n}^{\infty} e^{-\lambda t} \frac{(\lambda t)^j}{j!}$$

Conditional Distribution of the Arrival Times

Given that exactly one event has taken place in the interval $[0, t]$, determine the distribution of the time at which it occurred.

$$\begin{aligned}\Pr(X_1 < s | N(t) = 1) &= \frac{\Pr((N(s)=1) \cap (N(t)-N(s)=0))}{\Pr(N(t)=1)} \\ &= \frac{e^{-\lambda s} \lambda s e^{-\lambda(t-s)}}{e^{-\lambda t} \lambda t} = \frac{s}{t}\end{aligned}$$

That is, the time of the event is uniformly distributed in $[0, t]$.

If Y_1, \dots, Y_n are n random variables, then $Y_{(1)}, \dots, Y_{(n)}$ are the *order statistics* corresponding to Y_1, Y_2, \dots, Y_n if $Y_{(k)}$ is the k th smallest value among Y_1, \dots, Y_n , $k = 1, \dots, n$.

Theorem 1. *Given that $N(t) = n$, the n arrival times S_1, \dots, S_n have the same distribution as the order statistics corresponding to n independent random variables uniformly distributed on the interval $(0, t)$.*